

AN UPPER BOUND FOR THE CSISZÁR f -DIVERGENCE IN TERMS OF THE VARIATIONAL DISTANCE AND APPLICATIONS

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ABSTRACT. In this paper we point out an upper bound for the Csiszár f -divergence of two discrete random variables in terms of the generalized r -variational distance. Some particular cases for Kullback-Leibler distance, Hellinger and triangular discrimination, χ^2 -distance, etc.. are considered.

1. INTRODUCTION

Given a convex function $f : [0, \infty) \rightarrow \mathbb{R}$, the f -divergence functional

$$(1.1) \quad I_f(p, q) = \sum_{i=1}^n q_i f\left(\frac{p_i}{q_i}\right),$$

was introduced by Csiszár [1]-[2] as a generalized measure of information, a “distance function” on the set of probability distribution \mathbb{P}^n . The restriction here to discrete distributions is only for convenience, similar results hold for general distributions. As in Csiszár [1]-[2], we interpret undefined expressions by

$$\begin{aligned} f(0) &= \lim_{t \rightarrow 0^+} f(t), \quad 0f\left(\frac{0}{0}\right) = 0, \\ 0f\left(\frac{a}{0}\right) &= \lim_{\varepsilon \rightarrow 0^+} \varepsilon f\left(\frac{a}{\varepsilon}\right) = a \lim_{t \rightarrow \infty} \frac{f(t)}{t}, \quad a > 0. \end{aligned}$$

The following results (Theorems 1 and 2, and Corollary 1) were essentially given by Csiszár and Körner [3].

Theorem 1. (*Joint Convexity*) *If $f : [0, \infty) \rightarrow \mathbb{R}$ is convex, then $I_f(p, q)$ is jointly convex in p and q .*

Theorem 2. (*Jensen’s inequality*) *Let $f : [0, \infty) \rightarrow \mathbb{R}$ be convex. Then for any $p, q \in \mathbb{R}_+^n$ with $P_n := \sum_{i=1}^n p_i > 0$, $Q_n := \sum_{i=1}^n q_i > 0$, we have the inequality*

$$(1.2) \quad I_f(p, q) \geq Q_n f\left(\frac{P_n}{Q_n}\right).$$

If f is strictly convex, equality holds in (1.2) iff

$$(1.3) \quad \frac{p_1}{q_1} = \frac{p_2}{q_2} = \dots = \frac{p_n}{q_n}.$$

It is natural to consider the following corollary.

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Corollary 1. (Nonnegativity) Let $f : [0, \infty) \rightarrow \mathbb{R}$ be convex and normalised, i.e.,

$$(1.4) \quad f(1) = 0.$$

Then for any $p, q \in \mathbb{R}_+^n$ with $P_n = Q_n$, we have the inequality

$$(1.5) \quad I_f(p, q) \geq 0.$$

If f is strictly convex, equality holds in (1.5) iff

$$(1.6) \quad p_i = q_i \text{ for all } i \in \{1, \dots, n\}.$$

In particular, if p, q are probability vectors, then Corollary 1 shows that, for strictly convex and normalized $f : [0, \infty) \rightarrow \mathbb{R}$ that

$$(1.7) \quad I_f(p, q) \geq 0 \text{ and } I_f(p, q) = 0 \text{ iff } p = q.$$

We now give some more examples of divergence measures in Information Theory which are particular cases of Csiszár f -divergences.

1. **Kullback-Leibler distance** ([12]). The *Kullback-Leibler distance* $D(\cdot, \cdot)$ is defined by

$$(1.8) \quad D(p, q) := \sum_{i=1}^n p_i \log \left(\frac{p_i}{q_i} \right).$$

If we choose $f(t) = t \ln t$, $t > 0$, then obviously

$$(1.9) \quad I_f(p, q) = D(p, q).$$

2. **Variational distance** (l_1 -distance). The *variational distance* $V(\cdot, \cdot)$ is defined by

$$(1.10) \quad V(p, q) := \sum_{i=1}^n |p_i - q_i|.$$

If we choose $f(t) = |t - 1|$, $t \in \mathbb{R}_+$, then we have

$$(1.11) \quad I_f(p, q) = V(p, q).$$

3. **Hellinger discrimination** ([13]). The *Hellinger discrimination* $h^2(\cdot, \cdot)$ is defined by

$$(1.12) \quad h^2(p, q) := \frac{1}{2} \sum_{i=1}^n (\sqrt{p_i} - \sqrt{q_i})^2.$$

It is obvious that if $f(t) = \frac{1}{2} (\sqrt{t} - 1)^2$, then

$$(1.13) \quad I_f(p, q) = h^2(p, q).$$

4. **Triangular discrimination** ([24]). We define *triangular discrimination* between p and q by

$$(1.14) \quad \Delta(p, q) = \sum_{i=1}^n \frac{|p_i - q_i|^2}{p_i + q_i}.$$

It is obvious that if $f(t) = \frac{(t-1)^2}{t+1}$, $t \in (0, \infty)$, then

$$(1.15) \quad I_f(p, q) = \Delta(p, q).$$

5. χ^2 -**distance**. We define the χ^2 -*distance* (chi-square distance) by

$$(1.16) \quad D_{\chi^2}(p, q) := \sum_{i=1}^n \frac{(p_i - q_i)^2}{q_i}.$$

It is clear that if $f(t) = (t - 1)^2$, $t \in [0, \infty)$, then

$$(1.17) \quad I_f(p, q) = D_{\chi^2}(p, q)$$

6. **Rényi α -order entropy** ([14]). The α -*order entropy* ($\alpha > 1$) is defined by

$$(1.18) \quad R_\alpha(p, q) := \sum_{i=1}^n p_i^\alpha q_i^{1-\alpha}.$$

It is obvious that if $f(t) = t^\alpha$ ($t \in (0, \infty)$), then

$$(1.19) \quad I_f(p, q) = R_\alpha(p, q).$$

For other examples of divergence measures, see the paper [22] by J.N. Kapur, where further references are given.

2. AN UPPER BOUND FOR CSISZÁR f -DIVERGENCE

In the recent paper [28], the author proved the following inequality for Csiszár f -divergence:

Theorem 3. *Let $\Phi : \mathbb{R}_+ \rightarrow \mathbb{R}$ be differentiable convex. Then for all $p, q \in \mathbb{R}_+^n$, we have the inequality:*

$$(2.1) \quad \Phi'(1)(P_n - Q_n) \leq I_\Phi(p, q) - Q_n \Phi(1) \leq I_{\Phi'}\left(\frac{p^2}{q}, p\right) - I_{\Phi'}(p, q),$$

where $P_n := \sum_{i=1}^n p_i > 0$, $Q_n := \sum_{i=1}^n q_i > 0$ and $\Phi' : (0, \infty) \rightarrow \mathbb{R}$ is the derivative of Φ .

If Φ is strictly convex and $p_i, q_i > 0$ ($i = 1, \dots, n$), then the equality holds in (2.1) iff $p = q$,

If we assume that $P_n = Q_n$ and Φ is normalised, then we obtain the simpler

$$(2.2) \quad 0 \leq I_\Phi(p, q) \leq I_{\Phi'}\left(\frac{p^2}{q}, p\right) - I_{\Phi'}(p, q).$$

Applications for particular divergences which are instances of Csiszár f -divergence were also given.

A similar result of the above theorem has been presented in another paper by the author [29].

Theorem 4. *Let Φ, p, q be as in Theorem 3. Then we have the inequality*

$$(2.3) \quad 0 \leq I_\Phi(p, q) - Q_n \Phi\left(\frac{P_n}{Q_n}\right) \leq I_{\Phi'}\left(\frac{p^2}{q}, p\right) - \frac{P_n}{Q_n} I_{\Phi'}(p, q).$$

If Φ is strictly convex and $p_i, q_i > 0$ ($i = 1, \dots, n$), then the equality holds in (2.3) iff $\frac{p_1}{q_1} = \dots = \frac{p_n}{q_n}$.

Obviously, if $P_n = Q_n$ and Φ is normalised, then (2.3) becomes (2.2).

The following result concerning an upper and a lower bound for the Csiszár f -divergence in terms of the Kullback-Leibler distance $D(p, q)$ holds.

As in [30], we will say that the mapping $f : C \subset \mathbb{R} \rightarrow \mathbb{R}$, where C is an interval (in [30], the definition was considered in general normed spaces), is

- (i) α -lower convex on C if $f - \frac{\alpha}{2} \cdot |\cdot|^2$ is convex on C ;
- (ii) β -upper convex on C if $\frac{\beta}{2} \cdot |\cdot|^2 - f$ is convex on C ;
- (iii) (m, M) -convex on C (with $m \leq M$) if it is both m -lower convex and M -upper convex.

In [30], amongst others, the author has proved the following result for Csiszár f -divergence.

Theorem 5. *Let $\Phi : \mathbb{R}_+ \rightarrow \mathbb{R}$ and $p, q \in \mathbb{R}_+^n$ with $P_n = Q_n$.*

(i) *If Φ is α -lower convex on \mathbb{R}_+ , then we have the inequality*

$$(2.4) \quad \frac{\alpha}{2} \cdot D_{\chi^2}(p, q) \leq I_{\Phi}(p, q) - Q_n \Phi(1).$$

(ii) *If Φ is β -upper convex on \mathbb{R}_+ , then we have the inequality*

$$(2.5) \quad I_{\Phi}(p, q) - Q_n \Phi(1) \leq \frac{\beta}{2} \cdot D_{\chi^2}(p, q).$$

(iii) *If Φ is (m, M) -convex on \mathbb{R}_+ , then we have the following sandwich inequality*

$$(2.6) \quad \frac{m}{2} \cdot D_{\chi^2}(p, q) \leq I_{\Phi}(p, q) - Q_n \Phi(1) \leq \frac{M}{2} \cdot D_{\chi^2}(p, q),$$

where $D_{\chi^2}(\cdot, \cdot)$ is the χ^2 -divergence.

Of course, if Φ is normalised, i.e., $\Phi(1) = 0$ and p, q are probability distributions, then we get the simpler inequalities:

$$(2.7) \quad \frac{\alpha}{2} \cdot D_{\chi^2}(p, q) \leq I_{\Phi}(p, q), \quad I_{\Phi}(p, q) \leq \frac{\beta}{2} \cdot D_{\chi^2}(p, q)$$

and

$$(2.8) \quad \frac{m}{2} \cdot D_{\chi^2}(p, q) \leq I_{\Phi}(p, q) \leq \frac{M}{2} \cdot D_{\chi^2}(p, q).$$

In [30], some applications for particular instances of Csiszár f -divergences were also given.

Define the generalised r -variational distance by

$$(2.9) \quad V_r(p, q) := \sum_{i=1}^n q_i^{1-r} |p_i - q_i|^r,$$

where p, q are probability distributions and $r \in (0, 1]$. Note that for $r = 1$, we recapture the usual variational distance (or l_1 -distance).

The following theorem holds.

Theorem 6. *Let $f : [0, \infty) \rightarrow \mathbb{R}$ be a convex mapping and such that $f(1) = 0$, i.e., f is normalised, and f is of r -Hölder type on $[r, R]$, i.e.,*

$$(2.10) \quad |f(x) - f(y)| \leq H |x - y|^r \quad \text{for all } x, y \in [r, R].$$

Then we have the inequality

$$(2.11) \quad 0 \leq I_f(p, q) \leq H V_r(p, q).$$

Proof. We choose in (2.10) $x = \frac{p_i}{q_i}$, $y = 1$ ($i = 1, \dots, n$) to get

$$(2.12) \quad \left| f\left(\frac{p_i}{q_i}\right) - f(1) \right| \leq H \left| \frac{p_i}{q_i} - 1 \right|^r,$$

for all $i \in \{1, \dots, n\}$.

If we multiply (2.12) by q_i , sum the obtained inequalities and use the generalised triangle inequality, we obtain

$$\begin{aligned} 0 &\leq I_f(p, q) = \sum_{i=1}^n q_i f\left(\frac{p_i}{q_i}\right) = \left| \sum_{i=1}^n q_i f\left(\frac{p_i}{q_i}\right) - \sum_{i=1}^n q_i f(1) \right| \\ &\leq H \sum_{i=1}^n q_i \left| \frac{p_i}{q_i} - 1 \right|^r = HV_r(p, q) \end{aligned}$$

and the inequality (2.11) is proved. ■

Remark 1. *If we assume that f is convex, normalised and L -Lipschitzian on $[r, R]$, i.e., $r = 1$ and $H = L$, then we have the inequality*

$$(2.13) \quad 0 \leq I_f(p, q) \leq LV_r(p, q),$$

where $V(p, q)$ is the usual variational distance.

The practical result is embodied in the following corollary.

Corollary 2. *If the mapping $f : [r, R] \rightarrow \mathbb{R}$ is convex, normalised, absolutely continuous on $[r, R]$ and $f' \in L_\infty[r, R]$, i.e., $\|f'\|_\infty := \operatorname{ess\,sup}_{t \in [r, R]} |f'(t)| < \infty$,*

then we have the inequality

$$(2.14) \quad 0 \leq I_f(p, q) \leq \|f'\|_\infty V_r(p, q).$$

The following theorem holds.

Theorem 7. *Assume that f is as in Theorem 6. If $p^{(j)}, q^{(j)}$ ($j = 1, 2$) are probability distributions satisfying the condition*

$$r \leq \frac{p_i^{(j)}}{q_i^{(j)}} \leq R \quad \text{for all } i \in \{1, \dots, n\} \quad \text{and } j \in \{1, 2\},$$

then obviously

$$r \leq \frac{\lambda p_i^{(1)} + (1 - \lambda) p_i^{(2)}}{\lambda q_i^{(1)} + (1 - \lambda) q_i^{(2)}} \leq R \quad \text{for all } i \in \{1, \dots, n\} \quad \text{and } \lambda \in [0, 1]$$

and we have the inequality

$$\begin{aligned} (2.15) \quad 0 &\leq I_f\left(\lambda p^{(1)} + (1 - \lambda) p^{(2)}, \lambda q^{(1)} + (1 - \lambda) q^{(2)}\right) \\ &\quad - \lambda I_f\left(p^{(1)}, q^{(1)}\right) - (1 - \lambda) I_f\left(p^{(2)}, q^{(2)}\right) \\ &\leq H \lambda^r (1 - \lambda)^r \sum_{i=1}^n \frac{\left| \det \begin{bmatrix} p_i^{(1)} & p_i^{(2)} \\ q_i^{(1)} & q_i^{(2)} \end{bmatrix} \right|^r}{\left[\lambda q_i^{(1)} + (1 - \lambda) q_i^{(2)} \right]^r} \\ &\quad \times \left[\lambda^{1-r} \left[q_i^{(1)} \right]^{1-r} + (1 - \lambda)^{1-r} \left[q_i^{(2)} \right]^{1-r} \right] \end{aligned}$$

for all $\lambda \in [0, 1]$.

Proof. If we choose in (2.10)

$$x = \frac{\lambda p_i^{(1)} + (1 - \lambda) p_i^{(2)}}{\lambda q_i^{(1)} + (1 - \lambda) q_i^{(2)}} \quad \text{and} \quad y = \frac{p_i^{(1)}}{q_i^{(1)}} \quad (i = 1, \dots, n)$$

we get

$$\begin{aligned} (2.16) \quad & \left| f \left(\frac{\lambda p_i^{(1)} + (1 - \lambda) p_i^{(2)}}{\lambda q_i^{(1)} + (1 - \lambda) q_i^{(2)}} \right) - f \left(\frac{p_i^{(1)}}{q_i^{(1)}} \right) \right| \\ & \leq H \left| \frac{\lambda p_i^{(1)} + (1 - \lambda) p_i^{(2)}}{\lambda q_i^{(1)} + (1 - \lambda) q_i^{(2)}} - \frac{p_i^{(1)}}{q_i^{(1)}} \right|^r \\ & = H \frac{|\lambda q_i^{(1)} p_i^{(1)} + (1 - \lambda) q_i^{(1)} p_i^{(2)} - \lambda p_i^{(1)} q_i^{(1)} - (1 - \lambda) p_i^{(1)} q_i^{(2)}|^r}{[q_i^{(1)}]^r [\lambda q_i^{(1)} + (1 - \lambda) q_i^{(2)}]^r} \\ & = \frac{H (1 - \lambda)^r |q_i^{(1)} p_i^{(2)} - p_i^{(1)} q_i^{(2)}|^r}{[q_i^{(1)}]^r [\lambda q_i^{(1)} + (1 - \lambda) q_i^{(2)}]^r}. \end{aligned}$$

If we multiply (2.16) by $\lambda q_i^{(1)}$, we obtain

$$\begin{aligned} (2.17) \quad & \left| \lambda q_i^{(1)} f \left(\frac{\lambda p_i^{(1)} + (1 - \lambda) p_i^{(2)}}{\lambda q_i^{(1)} + (1 - \lambda) q_i^{(2)}} \right) - \lambda q_i^{(1)} f \left(\frac{p_i^{(1)}}{q_i^{(1)}} \right) \right| \\ & \leq \frac{H \lambda (1 - \lambda)^r [q_i^{(1)}]^{1-r} |q_i^{(1)} p_i^{(2)} - p_i^{(1)} q_i^{(2)}|^r}{[\lambda q_i^{(1)} + (1 - \lambda) q_i^{(2)}]^r} \end{aligned}$$

for all $i \in \{1, \dots, n\}$ and $\lambda \in [0, 1]$.

If in (2.10) we choose x as above but

$$y = \frac{p_i^{(2)}}{q_i^{(2)}} \quad (i = 1, \dots, n),$$

we get

$$\begin{aligned} (2.18) \quad & \left| f \left(\frac{\lambda p_i^{(1)} + (1 - \lambda) p_i^{(2)}}{\lambda q_i^{(1)} + (1 - \lambda) q_i^{(2)}} \right) - f \left(\frac{p_i^{(2)}}{q_i^{(2)}} \right) \right| \\ & \leq H \left| \frac{\lambda p_i^{(1)} + (1 - \lambda) p_i^{(2)}}{\lambda q_i^{(1)} + (1 - \lambda) q_i^{(2)}} - \frac{p_i^{(2)}}{q_i^{(2)}} \right|^r \\ & = H \frac{|\lambda p_i^{(1)} q_i^{(2)} + (1 - \lambda) p_i^{(2)} q_i^{(2)} - \lambda p_i^{(2)} q_i^{(1)} - (1 - \lambda) q_i^{(2)} p_i^{(2)}|^r}{[\lambda q_i^{(1)} + (1 - \lambda) q_i^{(2)}]^r [q_i^{(2)}]^r} \\ & = \frac{H (1 - \lambda)^r |p_i^{(1)} q_i^{(2)} - p_i^{(2)} q_i^{(1)}|^r}{[q_i^{(2)}]^r [\lambda q_i^{(1)} + (1 - \lambda) q_i^{(2)}]^r}. \end{aligned}$$

If we multiply (2.18) by $(1 - \lambda) q_i^{(2)}$, we obtain

$$(2.19) \quad \left| (1 - \lambda) q_i^{(2)} f \left(\frac{\lambda p_i^{(1)} + (1 - \lambda) p_i^{(2)}}{\lambda q_i^{(1)} + (1 - \lambda) q_i^{(2)}} \right) - (1 - \lambda) q_i^{(2)} f \left(\frac{p_i^{(2)}}{q_i^{(2)}} \right) \right| \\ \leq \frac{H \lambda^r (1 - \lambda) \left[q_i^{(2)} \right]^{1-r} \left| q_i^{(1)} p_i^{(2)} - p_i^{(1)} q_i^{(2)} \right|^r}{\left[\lambda q_i^{(1)} + (1 - \lambda) q_i^{(2)} \right]^r}$$

for all $i \in \{1, \dots, n\}$ and $\lambda \in [0, 1]$.

If we add (2.13) and (2.19) and use the triangle inequality, we get

$$(2.20) \quad \left| \left[\lambda q_i^{(1)} + (1 - \lambda) q_i^{(2)} \right] f \left(\frac{\lambda p_i^{(1)} + (1 - \lambda) p_i^{(2)}}{\lambda q_i^{(1)} + (1 - \lambda) q_i^{(2)}} \right) - \lambda q_i^{(1)} f \left(\frac{p_i^{(1)}}{q_i^{(1)}} \right) - (1 - \lambda) q_i^{(2)} f \left(\frac{p_i^{(2)}}{q_i^{(2)}} \right) \right| \\ \leq \frac{H \lambda^r (1 - \lambda)^r \left| \det \begin{bmatrix} p_i^{(1)} & p_i^{(2)} \\ q_i^{(1)} & q_i^{(2)} \end{bmatrix} \right|^r}{\left[\lambda q_i^{(1)} + (1 - \lambda) q_i^{(2)} \right]^r} \left[\lambda^{1-r} \left[q_i^{(1)} \right]^{1-r} + (1 - \lambda)^{1-r} \left[q_i^{(2)} \right]^{1-r} \right]$$

for all $i \in \{1, \dots, n\}$ and $\lambda \in [0, 1]$.

If we sum (2.20) over i from 1 to n and use the generalised triangle inequality, we obtain the desired inequality (2.15). ■

Remark 2. *If we assume that f is L -Lipschitzian, then the inequality (2.15) becomes*

$$(2.21) \quad 0 \leq I_f \left(\lambda p^{(1)} + (1 - \lambda) p^{(2)}, \lambda q^{(1)} + (1 - \lambda) q^{(2)} \right) - \lambda I_f \left(p^{(1)}, q^{(1)} \right) - (1 - \lambda) I_f \left(p^{(2)}, q^{(2)} \right) \\ \leq L \lambda (1 - \lambda) \sum_{i=1}^n \frac{\left| \det \begin{bmatrix} p_i^{(1)} & p_i^{(2)} \\ q_i^{(1)} & q_i^{(2)} \end{bmatrix} \right|}{\lambda q_i^{(1)} + (1 - \lambda) q_i^{(2)}},$$

for all $\lambda \in [0, 1]$.

3. SOME PARTICULAR CASES

Using the inequality (2.14), i.e.,

$$(3.1) \quad 0 \leq I_f(p, q) \leq \|f'\|_\infty V(p, q),$$

provided that f is absolutely continuous on $[r, R]$ and $f' \in L_\infty[r, R]$, $\|f'\|_\infty := \operatorname{ess\,sup}_{t \in [r, R]} |f'(t)|$, we shall be able to point out a number of interesting inequalities

between different divergence measures in Information Theory.

Proposition 1. *Let p, q be two probability distributions with the property that*

$$(3.2) \quad 0 < r \leq \frac{p_i}{q_i} =: r_i < R < \infty \quad \text{for all } i \in \{1, \dots, n\}.$$

Then we have the inequality

$$(3.3) \quad 0 \leq D(p, q) \leq \begin{cases} \left[\ln \sqrt{\frac{R}{r}} + \left| 1 + \ln \sqrt{rR} \right| \right] V(p, q) & \text{if } 0 < r \leq e^{-1} \\ (1 + \ln R) V(p, q) & \text{if } e^{-1} < r < 1 \end{cases}.$$

Proof. Consider the mapping $f : (0, \infty) \rightarrow \mathbb{R}$, $f(t) = t \ln t$. Then

$$f'(t) = \ln(et).$$

Therefore

$$\|f'\|_\infty = \sup_{t \in [r, R]} |f'(t)| = \max\{|\ln(er)|, \ln(eR)\}.$$

Now, observe that

1. If $0 < r \leq e^{-1}$, then $|\ln(er)| = \ln(er) = -1 - \ln r$ and then

$$\begin{aligned} \max\{|\ln(er)|, \ln(eR)\} &= \frac{-1 - \ln r + \ln R + 1 + |\ln r + 1 + 1 + \ln R|}{2} \\ &= \ln\left(\frac{R}{2}\right)^{\frac{1}{2}} + \left|1 + \ln \sqrt{rR}\right|. \end{aligned}$$

2. If $e^{-1} < r < 1$, then $|\ln(er)| = 1 + \ln r$ and then

$$\max\{|\ln(er)|, \ln(eR)\} = \max\{\ln(er), \ln(eR)\} = 1 + \ln R$$

and the proposition is proved. ■

We also have the following proposition.

Proposition 2. *Let p, q be as in Proposition 1. Then we have the inequality:*

$$(3.4) \quad 0 \leq D(q, p) \leq \frac{1}{r} V(p, q).$$

Proof. Consider the mapping $f(t) = -\ln t$, $t \in (0, \infty)$. Then

$$f'(t) = -\frac{1}{t} \quad \text{and} \quad \|f'\|_\infty = \frac{1}{r}.$$

Since

$$I_f(p, q) = -\sum_{i=1}^n q_i \ln\left(\frac{p_i}{q_i}\right) = \sum_{i=1}^n q_i \ln\left(\frac{q_i}{p_i}\right) = D(q, p),$$

the inequality (3.1) gives the desired result (3.4). ■

We point out now a bound for the χ^2 -divergence.

Proposition 3. *Let p, q be as above. Then*

$$(3.5) \quad 0 \leq D_{\chi^2}(p, q) \leq 2(R-1) V(p, q).$$

Proof. Consider the mapping $f : (0, \infty) \rightarrow \mathbb{R}$, $f(t) = (t-1)^2$. As $f'(t) = 2(t-1)$, it follows that $\|f'\|_\infty = 2(R-1)$. Using (3.1), we obtain (3.5). ■

The following result for Hellinger discrimination also holds.

Proposition 4. *Assume that the probability distributions p, q satisfy the condition (3.1). Then we have the inequality*

$$(3.6) \quad 0 \leq h^2(p, q) \leq \frac{1}{4} \left[\frac{\sqrt{R} - \sqrt{r}}{\sqrt{rR}} + \left| \frac{\sqrt{R} + \sqrt{r}}{\sqrt{rR}} - 2 \right| \right] V(p, q).$$

Proof. Consider the mapping $f : (0, \infty) \rightarrow \mathbb{R}$, $f(t) = \frac{1}{2}(\sqrt{t} - 1)^2$. Then

$$f'(t) = \frac{1}{2} - \frac{1}{2\sqrt{t}} \quad \text{and} \quad f''(t) = \frac{1}{4\sqrt{t^3}}, \quad t \in (0, \infty).$$

Then

$$\begin{aligned} \|f'\|_\infty &= \sup_{t \in [r, R]} |f'(t)| = \max \left\{ \left| \frac{\sqrt{r} - 1}{2\sqrt{r}} \right|, \left| \frac{\sqrt{R} - 1}{2\sqrt{R}} \right| \right\} \\ &= \max \left\{ \frac{1 - \sqrt{r}}{2\sqrt{r}}, \frac{\sqrt{R} - 1}{2\sqrt{R}} \right\} \\ &= \frac{1}{4} \left[\frac{\sqrt{R} - \sqrt{r}}{\sqrt{rR}} + \left| 2 - \frac{\sqrt{R} + \sqrt{r}}{\sqrt{rR}} \right| \right]. \end{aligned}$$

Using (3.1), we deduce (3.6). ■

Now, consider the Bhattacharyya distance which is defined by [27]

$$(3.7) \quad B(p, q) = \sum_{i=1}^n \sqrt{p_i q_i}.$$

The following proposition holds.

Proposition 5. *Assume that p, q are probability distributions. Then we have the inequality*

$$(3.8) \quad 0 \leq 1 - B(p, q) \leq V_{\frac{1}{2}}(p, q),$$

where $V_{\frac{1}{2}}(p, q) = \sum_{i=1}^n \sqrt{q_i(p_i - q_i)}$ is the $\frac{1}{2}$ -variational distance.

Proof. Consider the mapping $f : [0, \infty) \rightarrow [0, \infty)$ given by

$$f(t) = -\sqrt{t} + 1.$$

Obviously,

$$|f(x) - f(y)| = |\sqrt{x} - \sqrt{y}| \leq \sqrt{|x - y|}, \quad \text{for all } x, y \in [0, \infty),$$

which shows that f is of $\frac{1}{2}$ -Hölder type with the constant $H = 1$.

Applying Theorem 6, we deduce (3.8). ■

Another inequality for Bhattacharyya distance in terms of the variational distance V is embodied in the following proposition.

Proposition 6. *Assume that the probability distributions p, q satisfy the condition:*

$$(3.9) \quad 0 < r \leq \frac{p_i}{q_i} \quad \text{for all } i \in \{1, \dots, n\}.$$

Then we have the inequality

$$(3.10) \quad 0 \leq 1 - B(p, q) \leq \frac{1}{2\sqrt{r}} V(p, q).$$

Proof. For the mapping $f(t) = -\sqrt{t} + 1$, we have $f'(t) = -\frac{1}{2\sqrt{t}}$ and $\|f'\|_\infty = \sup_{t \in [0, \infty)} \frac{1}{2\sqrt{t}}$. Applying (3.1), we deduce (3.11). ■

Now, we consider another useful divergence measure in Information Theory which is known as the *Harmonic distance*:

$$(3.11) \quad M(p, q) := \sum_{i=1}^n \frac{2p_i q_i}{p_i + q_i}.$$

Proposition 7. *Assume that the probability distributions p, q satisfy the condition (3.9). Then we have the inequality*

$$(3.12) \quad 0 \leq 1 - M(p, q) \leq \frac{2}{(r+1)^2} V(p, q).$$

Proof. Consider the mapping $f: [0, \infty) \rightarrow \mathbb{R}$, $f(t) = 1 - \frac{2t}{t+1}$. Obviously $f'(t) = -\frac{2}{(t+1)^2}$, $f''(t) = \frac{4}{(t+1)^3}$ and $\|f'\|_\infty = \sup_{t \in [r, \infty)} |f'(t)| = \frac{2}{(r+1)^2}$. As

$$I_f(p, q) = 1 - M(p, q),$$

then by (3.1) we deduce the desired inequality (3.12). ■

Another useful divergence measure is the Jeffreys J -divergence defined by [26]

$$J(p, q) = \sum_{i=1}^n (p_i - q_i) \ln \left(\frac{p_i}{q_i} \right).$$

We can state the following proposition as well.

Proposition 8. *Assume that the probability distributions p, q satisfy the condition:*

$$(3.13) \quad \frac{p_i}{q_i} \leq R < \infty, \quad (i = 1, \dots, n).$$

Then we have the inequality

$$(3.14) \quad 0 \leq J(p, q) \leq \left(\ln R - \frac{1}{R} + 1 \right) V(p, q).$$

Proof. Consider the mapping $f(t) = (t-1) \ln t$, $t > 0$. Then

$$\begin{aligned} f'(t) &= \ln t - \frac{1}{t} + 1; \quad t \in (0, \infty) \\ f''(t) &= \frac{t+1}{t^2}, \quad t \in (0, \infty) \end{aligned}$$

which shows that

$$\|f'\|_\infty = \sup_{t \in (0, R]} |f'(t)| = f'(R) = \ln R - \frac{1}{R} + 1.$$

As

$$I_f(p, q) = J(p, q),$$

then by (3.1) we deduce (3.14). ■

Finally, the following result for the triangular discrimination holds (for the definition of triangular discrimination, see [24]).

Proposition 9. *If p, q are such that the condition (3.13) holds, then*

$$(3.15) \quad 0 \leq \Delta(p, q) \leq \frac{(R-1)(R+3)}{(R+1)^2} V(p, q) \leq V(p, q).$$

The proof is obvious by (3.2) applied for the mapping $f(t) = \frac{(t-1)^2}{t+1}$, and we omit the details.

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